

Class Notes: Problem Set 5

EC337

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1 Some useful properties

1. Bilinearity and associativity of the Kronecker product:

For arbitrary conformable matrices A , B , and C and scalar k ,

$$A \otimes (B + C) = A \otimes B + A \otimes C,$$

$$(B + C) \otimes A = B \otimes A + C \otimes A,$$

$$(kA) \otimes B = A \otimes (kB) = k(A \otimes B),$$

$$(A \otimes B) \otimes C = A \otimes (B \otimes C),$$

$$A \otimes 0 = 0 \otimes A = 0.$$

2. Mixed-product property of the Kronecker product:

For arbitrary conformable matrices A , B , C , and D ,

$$(A \otimes B)(C \otimes D) = (AC) \otimes (BD).$$

3. Mixed Kronecker matrix-vector product:

For arbitrary conformable matrices A , B , and C ,

$$\text{vec}(ABC) = (C' \otimes A) \text{vec}(B).$$

4. Inverse of a Kronecker product:

For arbitrary nonsingular matrices A and B ,

$$(A \otimes B)^{-1} = A^{-1} \otimes B^{-1}.$$

5. Transpose of a Kronecker product:

For arbitrary matrices A and B ,

$$(A \otimes B)' = A' \otimes B'.$$

6. Determinant of a Kronecker product:

For arbitrary $n \times n$ matrix A and $k \times k$ matrix B ,

$$|A \otimes B| = |A|^k |B|^n.$$

7. Trace of a Kronecker product:

For arbitrary square matrices A and B ,

$$\text{tr}(A \otimes B) = \text{tr}(A) \text{tr}(B).$$

8. Vectorization is a unitary transformation (preserves the inner product):

For arbitrary $m \times n$ matrices A and B ,

$$\text{tr}(A'B) = \text{vec}(A)' \text{vec}(B).$$

9. Basic properties of the trace:

For arbitrary square matrices A and B and scalar k ,

$$\text{tr}(A + B) = \text{tr}(A) + \text{tr}(B),$$

$$\text{tr}(kA) = k \text{tr}(A),$$

$$\text{tr}(A') = \text{tr}(A),$$

$$\text{tr}(I_k) = k.$$

10. The trace is invariant under cyclic permutations (the cyclic property):

For arbitrary suitably conformable matrices A , B , C , and D ,

$$\text{tr}(ABCD) = \text{tr}(BCDA) = \text{tr}(CDAB) = \text{tr}(DABC).$$

11. (Some) basic properties of the determinant:

For arbitrary $n \times n$ matrices A and B and scalar k ,

$$|I_k| = 1,$$

$$|kA| = k^n |A|,$$

$$|A'| = |A|,$$

$$|A^{-1}| = |A|^{-1} \text{ (if } |A| \neq 0),$$

$$|AB| = |A| |B|.$$

12. Determinant of a triangular matrix:

For arbitrary triangular matrix A with diagonal entries $\{a_{ii}\}_{i=1}^n$,

$$|A| = \prod_{i=1}^n a_{ii}.$$

Moreover,

$$|A| = 0 \iff \exists i \in \{1, \dots, n\} : a_{ii} = 0,$$

and

$$a_{ii} = 1 \ \forall i \in \{1, \dots, n\} \implies |A| = 1.$$

13. Basic properties of triangular matrices:

Let $A, B \in \Delta$ be two arbitrary suitably conformable lower (upper) triangular matrices and k a scalar, where Δ is the set of all lower (upper) triangular matrices and ∇ is the set of all upper (lower) triangular matrices. Then,

$$kA \in \Delta,$$

$$A + B \in \Delta,$$

$$AB \in \Delta,$$

$$A^{-1} \in \Delta \text{ with diagonal } \{a_{ii}^{-1}\}_{i=1}^n \text{ (provided } |A| \neq 0 \text{)},$$

$$A' \in \nabla$$

2 Notation

Consider the following system of N equations in N endogenous variables and K exogenous variables for a sample of T observations $t \in \{1, \dots, T\}$

$$\tilde{a}_1^1 y_{1t} + \dots + \tilde{a}_N^1 y_{Nt} + \tilde{b}_1^1 x_{1t} + \dots + \tilde{b}_K^1 x_{Kt} = u_{1t} \tag{S.1}$$

$$\vdots$$

$$\tilde{a}_1^N y_{1t} + \dots + \tilde{a}_N^N y_{Nt} + \tilde{b}_1^N x_{1t} + \dots + \tilde{b}_K^N x_{Kt} = u_{Nt}. \tag{S.N}$$

Let

$$y_t \equiv \begin{pmatrix} y_{1t} \\ \vdots \\ y_{Nt} \end{pmatrix}_{N \times 1} \quad x_t \equiv \begin{pmatrix} x_{1t} \\ \vdots \\ x_{Kt} \end{pmatrix}_{K \times 1} \quad u_t \equiv \begin{pmatrix} u_{1t} \\ \vdots \\ u_{Nt} \end{pmatrix}_{N \times 1}$$

$$\tilde{A} \equiv \begin{bmatrix} \tilde{a}_1^1 & \cdots & \tilde{a}_N^1 \\ \vdots & \ddots & \vdots \\ \tilde{a}_1^N & \cdots & \tilde{a}_N^N \end{bmatrix}_{N \times N} \quad \tilde{B} \equiv \begin{bmatrix} \tilde{b}_1^1 & \cdots & \tilde{b}_K^1 \\ \vdots & \ddots & \vdots \\ \tilde{b}_1^K & \cdots & \tilde{b}_K^K \end{bmatrix}_{N \times K}$$

Then, we can compactly write the system for observation t as

$$\tilde{A}y_t + \tilde{B}x_t = u_t \quad (1)$$

and, transposing and stacking the system for the T observations, we obtain the matrix equation

$$Y\tilde{A}' + X\tilde{B}' = U, \quad (2)$$

where

$$Y \equiv \begin{bmatrix} y_{11} & \cdots & y_{N1} \\ \vdots & \ddots & \vdots \\ y_{1T} & \cdots & y_{NT} \end{bmatrix}_{T \times N} = \begin{bmatrix} y'_1 \\ \vdots \\ y'_T \end{bmatrix}$$

$$X \equiv \begin{bmatrix} x_{11} & \cdots & x_{K1} \\ \vdots & \ddots & \vdots \\ x_{1T} & \cdots & x_{KT} \end{bmatrix}_{T \times K} = \begin{bmatrix} x'_1 \\ \vdots \\ x'_T \end{bmatrix}$$

$$U \equiv \begin{bmatrix} u_{11} & \cdots & u_{N1} \\ \vdots & \ddots & \vdots \\ u_{1T} & \cdots & u_{NT} \end{bmatrix}_{T \times N} = \begin{bmatrix} u'_1 \\ \vdots \\ u'_T \end{bmatrix}$$

As discussed in the lecture notes, the system is not identified without further restrictions on the structural parameters

$$\tilde{\delta} \equiv \begin{bmatrix} \text{vec}(\tilde{A}) \\ \text{vec}(\tilde{B}) \end{bmatrix}_{N(N+K) \times 1}.$$

From now on we will suppose the system is identified through the imposition of

- (i) N normalizations $\tilde{a}_i^i = 1$, $i \in \{1, \dots, N\}$
- (ii) At least $N(N-1)$ exclusion restrictions of the form $\tilde{a}_j^i = 0$ or $\tilde{b}_k^i = 0$, $i, j \in \{1, \dots, N\}$. $i \neq j$, $k \in \{1, \dots, K\}$.

For $i \in \{1, \dots, N\}$, define:

- $y_t^{(i)}$: $N_i \times 1$ subvector of y_t , containing $N_i \leq N - 1$ of the $N - 1$ endogenous variables in y_t excluding y_{it} —i.e., the endogenous variable on the LHS of equation $(S.i)$. The endogenous variables y_{jt} (where $j \neq i$) included in $y_t^{(i)}$ are those with $\tilde{a}_j^i \neq 0$.

- $a^{(i)}$: $N_i \times 1$ subvector of

$$a^i \equiv \begin{pmatrix} a_1^i \\ \vdots \\ a_N^i \end{pmatrix}_{N \times 1}$$

corresponding to the coefficients on the endogenous variables y_{jt} in $y_t^{(i)}$, i.e., those with $a_j^i \neq 0$, where $a_j^i \equiv -\tilde{a}_j^i$ (for $i \neq j$).

- $x_t^{(i)}$: $K_i \times 1$ subvector of x_t , containing $K_i \leq K$ of the K exogenous variables in x_t . These are the exogenous variables on the RHS of equation $(S.i)$. The exogenous variables x_{kt} included in $x_t^{(i)}$ are those with $\tilde{b}_k^i \neq 0$.

- $b^{(i)}$: $K_i \times 1$ subvector of

$$b^i \equiv \begin{pmatrix} b_1^i \\ \vdots \\ b_K^i \end{pmatrix}_{K \times 1}$$

corresponding to the coefficients on the exogenous variables x_{kt} in $x_t^{(i)}$, i.e., those with $b_k^i \neq 0$, where $b_k^i \equiv -\tilde{b}_k^i$.

Then, we can write the system after imposing the normalizations and exclusion restrictions as

$$y_{1t} = y_t^{(1)'} a^{(1)} + x_t^{(1)'} b^{(1)} + u_{1t} \quad (S.1)$$

\vdots

$$y_{Nt} = y_t^{(N)'} a^{(N)} + x_t^{(N)'} b^{(N)} + u_{Nt}. \quad (S.N)$$

Now, let

$$z_t^{(i)} \equiv \begin{pmatrix} y_t^{(i)} \\ x_t^{(i)} \end{pmatrix}_{(N_i + K_i) \times 1}$$

and

$$\delta^{(i)} \equiv \begin{pmatrix} a^{(i)} \\ b^{(i)} \end{pmatrix}_{(N_i + K_i) \times 1}$$

so that equation $(S.i)$ can be written as

$$y_{it} = z_t^{(i)'} \delta^{(i)} + u_{it}$$

for $i \in \{1, \dots, N\}$. Stacking equation $(S.i)$ for the T observations $t \in \{1, \dots, T\}$, we obtain the matrix form

$$Y_i = Z^{(i)}\delta^{(i)} + U_i,$$

where Y_i and U_i are the i^{th} columns of matrices Y and U , respectively,

$$Z^{(i)} \equiv \begin{bmatrix} Y^{(i)} & X^{(i)} \end{bmatrix}_{T \times (N_i + K_i)}$$

$$Y^{(i)} \equiv \begin{bmatrix} y_1^{(i)'} \\ \vdots \\ y_T^{(i)'} \end{bmatrix}_{T \times N_i}$$

$$X^{(i)} \equiv \begin{bmatrix} x_1^{(i)'} \\ \vdots \\ x_T^{(i)'} \end{bmatrix}_{T \times K_i}.$$

3 Question 1

3.1 General derivations

3.1.1 2SLS estimator

First, consider equation $(S.i)$ alone, where $i \in \{1, \dots, N\}$. The goal here is estimation of $\delta^{(i)}$ and, as stated above, identification is assumed. Since we are interested in the i^{th} column of Y , Y_i , we need excluded instruments for $Y^{(i)} — X^{(i)}$ act as included instruments for themselves.

Solving for Y in equation (2), we obtain the reduced form

$$Y = X\Pi' + V, \tag{3}$$

where

$$\Pi' \equiv -\tilde{B}'\tilde{A}'^{-1}$$

and

$$V \equiv U\tilde{A}'^{-1}.$$

Notice that exogeneity of X implies that

$$\begin{aligned} \mathbb{E}[X'V] &= \mathbb{E}[X'U]\tilde{A}'^{-1} \\ &= 0_{K \times N}, \end{aligned}$$

so $\widehat{Y} = X\widehat{\Pi}'$ and U —where $\widehat{\Pi}' = (X'X)^{-1}X'Y$ is the OLS estimator of Π' in multivariate regression (3)— are asymptotically uncorrelated since $\widehat{\Pi} \xrightarrow{p} \Pi$. Therefore, for the best linear predictor of $Y^{(i)}$, $\widehat{Y}^{(i)}$, is a valid excluded instrument, and the full matrix of instruments is

$$\widehat{Z}^{(i)} \equiv \begin{bmatrix} \widehat{Y}^{(i)} & X^{(i)} \end{bmatrix}_{T \times (N_i + K_i)}.$$

Note that $\widehat{Y}^{(i)}$ is a submatrix of

$$\begin{aligned} \widehat{Y} &= X\widehat{\Pi}' \\ &= X(X'X)^{-1}X'Y \\ &= X(X'X)^{-1}X' \begin{bmatrix} Y_1 & \dots & Y_N \end{bmatrix} \\ &= \begin{bmatrix} X(X'X)^{-1}X'Y_1 & \dots & X(X'X)^{-1}X'Y_N \end{bmatrix}, \end{aligned}$$

formed by the N_i columns corresponding to the variables in $y_t^{(i)}$, i.e.,

$$\widehat{Y}^{(i)} = X(X'X)^{-1}X'Y^{(i)}.$$

Moreover, since $X^{(i)} \in \text{Col}(X) \implies X(X'X)^{-1}X'X^{(i)} = X^{(i)}$,

$$\begin{aligned} \widehat{Y}^{(i)'}X^{(i)} &= \left(X(X'X)^{-1}X'Y^{(i)} \right)' X^{(i)} \\ &= Y^{(i)'}X(X'X)^{-1}X'X^{(i)} \\ &= Y^{(i)'}X^{(i)} \end{aligned}$$

and

$$\begin{aligned} \widehat{Y}^{(i)'}\widehat{Y}^{(i)} &= \left(X(X'X)^{-1}X'Y^{(i)} \right)' X(X'X)^{-1}X'Y^{(i)} \\ &= Y^{(i)'}X(X'X)^{-1}X'X(X'X)^{-1}X'Y^{(i)} \\ &= Y^{(i)'}X(X'X)^{-1}X'Y^{(i)} \\ &= \widehat{Y}^{(i)'}Y^{(i)} \\ &= Y^{(i)'}\widehat{Y}^{(i)}. \end{aligned}$$

Therefore,

$$\widehat{Z}^{(i)'}Z^{(i)} = \begin{bmatrix} \widehat{Y}^{(i)'} \\ X^{(i)'} \end{bmatrix} \begin{bmatrix} Y^{(i)} & X^{(i)} \end{bmatrix}$$

$$\begin{aligned}
&= \begin{bmatrix} \widehat{Y}^{(i)'} Y^{(i)} & \widehat{Y}^{(i)'} X^{(i)} \\ X^{(i)'} Y^{(i)} & X^{(i)'} X^{(i)} \end{bmatrix} \\
&= \begin{bmatrix} \widehat{Y}^{(i)'} \widehat{Y}^{(i)} & \widehat{Y}^{(i)'} X^{(i)} \\ X^{(i)'} \widehat{Y}^{(i)} & X^{(i)'} X^{(i)} \end{bmatrix} \\
&= \begin{bmatrix} \widehat{Y}^{(i)'} \\ X^{(i)'} \end{bmatrix} \begin{bmatrix} \widehat{Y}^{(i)} & X^{(i)} \end{bmatrix} \\
&= \widehat{Z}^{(i)'} \widehat{Z}^{(i)},
\end{aligned}$$

so the IV estimator is

$$\begin{aligned}
\widehat{\delta}_{2\text{SLS}}^{(i)} &= (\widehat{Z}^{(i)'} \widehat{Z}^{(i)})^{-1} \widehat{Z}^{(i)'} Y_i \\
&= (\widehat{Z}^{(i)'} \widehat{Z}^{(i)})^{-1} \widehat{Z}^{(i)'} Y_i,
\end{aligned}$$

i.e., the OLS estimator of the regression of Y_i on $\widehat{Y}^{(i)}$ and $X^{(i)}$.

Now, consider estimation of the full system. To this end, write the system in matrix form by stacking the N previously derived matrix equations

$$Y_i = Z^{(i)} \delta^{(i)} + U_i$$

for $i \in \{1, \dots, N\}$ as follows:

$$\begin{pmatrix} Y_1 \\ \vdots \\ Y_N \end{pmatrix}_{NT \times 1} = \begin{bmatrix} Z^{(1)} & \cdots & 0_{T \times (N_N + K_N)} \\ \vdots & \ddots & \vdots \\ 0_{T \times (N_1 + K_1)} & \cdots & Z^{(N)} \end{bmatrix}_{NT \times \left(\sum_{i=1}^N N_i + K_i \right)} \begin{pmatrix} \delta^{(1)} \\ \vdots \\ \delta^{(N)} \end{pmatrix}_{\left(\sum_{i=1}^N N_i + K_i \right) \times 1} + \begin{pmatrix} U_1 \\ \vdots \\ U_N \end{pmatrix}_{NT \times 1}$$

$$\Longleftrightarrow Y^* = Z^* \delta + U^*,$$

where Y^* , Z^* , δ , and U^* are defined in the obvious way.¹

Similarly, let

$$\widehat{Z}^* \equiv \begin{bmatrix} \widehat{Z}^{(1)} & \cdots & 0_{T \times (N_N + K_N)} \\ \vdots & \ddots & \vdots \\ 0_{T \times (N_1 + K_1)} & \cdots & \widehat{Z}^{(N)} \end{bmatrix}_{NT \times \left(\sum_{i=1}^N N_i + K_i \right)}$$

¹Note that $Y^* = \text{vec}(Y)$.

$$\begin{aligned}
&= \begin{bmatrix} X(X'X)^{-1}X'Z^{(1)} & \cdots & 0_{T \times (N_N + K_N)} \\ \vdots & \ddots & \vdots \\ 0_{T \times (N_1 + K_1)} & \cdots & X(X'X)^{-1}X'Z^{(N)} \end{bmatrix} \\
&= \underbrace{\begin{bmatrix} X(X'X)^{-1}X' & \cdots & 0_{T \times T} \\ \vdots & \ddots & \vdots \\ 0_{T \times T} & \cdots & X(X'X)^{-1}X' \end{bmatrix}}_{NT \times NT} \underbrace{\begin{bmatrix} Z^{(1)} & \cdots & 0_{T \times (N_N + K_N)} \\ \vdots & \ddots & \vdots \\ 0_{T \times (N_1 + K_1)} & \cdots & Z^{(N)} \end{bmatrix}}_{NT \times \left(\sum_{i=1}^N N_i + K_i\right)} \\
&= \left(I_N \otimes X(X'X)^{-1}X'\right)Z^* \\
&= \left(I_N I_N \otimes X((X'X)^{-1}X')\right)Z^* \\
&= \left(I_N \otimes X\right)\left(I_N \otimes (X'X)^{-1}X'\right)Z^* \\
&= \left(I_N \otimes X\right)\left(I_N I_N \otimes ((X'X)^{-1})X'\right)Z^* \\
&= \left(I_N \otimes X\right)\left(I_N \otimes (X'X)^{-1}\right)\left(I_N \otimes X'\right)Z^*.
\end{aligned}$$

Finally, following analogous arguments to those discussed above for the single equation case, the system-2SLS estimator is the OLS estimator of the regression of Y^* on \hat{Z}^* ,

$$\begin{aligned}
\hat{\delta}_{2SLS} &= (\hat{Z}^{*'}Z^*)^{-1}\hat{Z}^{*'}Y^* \\
&= (\hat{Z}^{*'}\hat{Z}^*)^{-1}\hat{Z}^{*'}Y^*.
\end{aligned}$$

3.1.2 3SLS estimator

Now, to obtain the 3SLS estimator—which amounts to a GLS-style transformation—, notice that the assumption that

$$u_t \stackrel{iid}{\sim} (0, \Sigma)$$

implies that, for $i, j \in \{1, \dots, N\}$,

$$\mathbb{E}[U_i U_j'] = \mathbb{E}\left[\begin{pmatrix} u_{i1} \\ \vdots \\ u_{iT} \end{pmatrix} \begin{pmatrix} u_{j1} & \cdots & u_{jT} \end{pmatrix}\right]$$

$$\begin{aligned}
&= \begin{bmatrix} \mathbb{E}[u_{i1}u_{j1}] & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \mathbb{E}[u_{iT}u_{jT}] \end{bmatrix} \\
&= \begin{bmatrix} \sigma_{ij} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \sigma_{ij} \end{bmatrix} \\
&= \sigma_{ij} I_T,
\end{aligned}$$

where

$$\Sigma \equiv \begin{bmatrix} \sigma_{11} & \cdots & \sigma_{1N} \\ \vdots & \ddots & \vdots \\ \sigma_{N1} & \cdots & \sigma_{NN} \end{bmatrix}.$$

Therefore,

$$\begin{aligned}
\mathbb{E}[U^*U^{*'}] &= \mathbb{E} \left[\begin{pmatrix} U_1 \\ \vdots \\ U_N \end{pmatrix} (U_1' \cdots U_N') \right] \\
&= \begin{bmatrix} \mathbb{E}[U_1U_1'] & \cdots & \mathbb{E}[U_1U_N'] \\ \vdots & \ddots & \vdots \\ \mathbb{E}[U_NU_1'] & \cdots & \mathbb{E}[U_NU_N'] \end{bmatrix} \\
&= \begin{bmatrix} \sigma_{11} I_T & \cdots & \sigma_{1N} I_T \\ \vdots & \ddots & \vdots \\ \sigma_{N1} I_T & \cdots & \sigma_{NN} I_T \end{bmatrix} \\
&= \begin{bmatrix} \sigma_{11} & \cdots & \sigma_{1N} \\ \vdots & \ddots & \vdots \\ \sigma_{N1} & \cdots & \sigma_{NN} \end{bmatrix} \otimes I_T \\
&= \Sigma \otimes I_T.
\end{aligned}$$

Next, applying the GLS-style transformation,

$$\left(\Sigma^{-1/2} \otimes I_T \right) Y^* = \left(\Sigma^{-1/2} \otimes I_T \right) Z^* \delta + \left(\Sigma^{-1/2} \otimes I_T \right) U^*$$

with

$$\begin{aligned}
\mathbb{E} \left[\left(\Sigma^{-1/2} \otimes I_T \right) U^* \left(\left(\Sigma^{-1/2} \otimes I_T \right) U^* \right)' \right] &= \left(\Sigma^{-1/2} \otimes I_T \right) \mathbb{E} [U^* U^{*'}] \left(\Sigma^{-1/2} \otimes I_T \right)' \\
&= \left(\Sigma^{-1/2} \otimes I_T \right) \left(\Sigma \otimes I_T \right) \left(\Sigma^{-1/2'} \otimes I_T' \right) \\
&= \left(\Sigma^{-1/2} \Sigma \otimes I_T I_T \right) \left(\Sigma^{-1/2} \otimes I_T \right) \\
&= \left(\Sigma^{-1/2} \Sigma \Sigma^{-1/2} \otimes I_T I_T I_T \right) \\
&= \left(I_N \otimes I_T \right) \\
&= I_{NT},
\end{aligned}$$

we obtain the 3SLS estimator

$$\begin{aligned}
\hat{\delta}_{3SLS} &= \left(\left(\left(\Sigma^{-1/2} \otimes I_T \right) \hat{Z}^* \right)' \left(\Sigma^{-1/2} \otimes I_T \right) \hat{Z}^* \right)^{-1} \left(\left(\Sigma^{-1/2} \otimes I_T \right) \hat{Z}^* \right)' \left(\Sigma^{-1/2} \otimes I_T \right) Y^* \\
&= \left(\hat{Z}^{*'} \left(\Sigma^{-1/2'} \otimes I_T' \right) \left(\Sigma^{-1/2} \otimes I_T \right) \hat{Z}^* \right)^{-1} \hat{Z}^{*'} \left(\Sigma^{-1/2'} \otimes I_T' \right) \left(\Sigma^{-1/2} \otimes I_T \right) Y^* \\
&= \left(\hat{Z}^{*'} \left(\Sigma^{-1/2} \Sigma^{-1/2} \otimes I_T I_T \right) \hat{Z}^* \right)^{-1} \hat{Z}^{*'} \left(\Sigma^{-1/2} \Sigma^{-1/2} \otimes I_T I_T \right) Y^* \\
&= \left(\hat{Z}^{*'} \left(\Sigma^{-1} \otimes I_T \right) \hat{Z}^* \right)^{-1} \hat{Z}^{*'} \left(\Sigma^{-1} \otimes I_T \right) Y^* \\
&= \left(Z^{*'} X (X' X)^{-1} X' \left(\Sigma^{-1} \otimes I_T \right) X (X' X)^{-1} X' Z^* \right)^{-1} Z^{*'} X (X' X)^{-1} X' \left(\Sigma^{-1} \otimes I_T \right) Y^* \\
&= \left(Z^{*'} \left(1 \otimes X (X' X)^{-1} X' \right) \left(\Sigma^{-1} \otimes I_T \right) X (X' X)^{-1} X' Z^* \right)^{-1} Z^{*'} \left(1 \otimes X (X' X)^{-1} X' \right) \left(\Sigma^{-1} \otimes I_T \right) Y^* \\
&= \left(Z^{*'} \left(\Sigma^{-1} \otimes X (X' X)^{-1} X' \right) X (X' X)^{-1} X' Z^* \right)^{-1} Z^{*'} \left(\Sigma^{-1} \otimes X (X' X)^{-1} X' \right) Y^* \\
&= \left(Z^{*'} \left(\Sigma^{-1} \otimes X (X' X)^{-1} X' \right) \left(1 \otimes X (X' X)^{-1} X' \right) Z^* \right)^{-1} Z^{*'} \left(\Sigma^{-1} \otimes X (X' X)^{-1} X' \right) Y^* \\
&= \left(Z^{*'} \left(\Sigma^{-1} \otimes X (X' X)^{-1} X' X (X' X)^{-1} X' \right) Z^* \right)^{-1} Z^{*'} \left(\Sigma^{-1} \otimes X (X' X)^{-1} X' \right) Y^*
\end{aligned}$$

$$\begin{aligned}
&= \left(Z^{*'} \left(\Sigma^{-1} \otimes X(X'X)^{-1}X' \right) Z^* \right)^{-1} Z^{*'} \left(\Sigma^{-1} \otimes X(X'X)^{-1}X' \right) Y^* \\
&= \left(Z^{*'} \left(I_N \Sigma^{-1} \otimes X(X'X)^{-1}X' \right) Z^* \right)^{-1} Z^{*'} \left(I_N \Sigma^{-1} \otimes X(X'X)^{-1}X' \right) Y^* \\
&= \left(Z^{*'} \left(I_N \otimes X \right) \left(\Sigma^{-1} \otimes (X'X)^{-1}X' \right) Z^* \right)^{-1} Z^{*'} \left(I_N \otimes X \right) \left(\Sigma^{-1} \otimes (X'X)^{-1}X' \right) Y^* \\
&= \left(Z^{*'} \left(I_N \otimes X \right) \left(\Sigma^{-1} I_N \otimes (X'X)^{-1}X' \right) Z^* \right)^{-1} Z^{*'} \left(I_N \otimes X \right) \left(\Sigma^{-1} I_N \otimes (X'X)^{-1}X' \right) Y^* \\
&= \left(Z^{*'} \left(I_N \otimes X \right) \left(\Sigma^{-1} \otimes (X'X)^{-1} \right) \left(I_N \otimes X' \right) Z^* \right)^{-1} Z^{*'} \left(I_N \otimes X \right) \left(\Sigma^{-1} \otimes (X'X)^{-1} \right) \left(I_N \otimes X' \right) Y^*.
\end{aligned}$$

3.2 Exactly identified system

Finally, notice that exact identification requires that $N_i + K_i = K \ \forall i \in \{1, \dots, N\}$. The necessary order condition for identification of equation $(S.i)$ is that the number of restrictions we impose on the $N + K$ parameters on the LHS (of the first representation discussed above) through the normalization $\tilde{a}_i^i = 1$ and the exclusion restrictions —i.e., $(N - N_i) + (K - K_i)$ — is at least N .² That is,

$$N - N_i + K - K_i \geq N \iff K - K_i \geq N_i,$$

which requires that the number of excluded exogenous variables appearing elsewhere in the system, $K - K_i$, is at least as large as the number of included endogenous variables, N_i . Since the equation is just identified, the order condition holds with equality and we get

$$K - K_i = N_i \iff N_i + K_i = K.$$

Therefore, block diagonal matrix

$$\left(I_N \otimes X' \right) Z^* = \underbrace{\begin{bmatrix} X' & \cdots & 0_{K \times T} \\ \vdots & \ddots & \vdots \\ 0_{K \times T} & \cdots & X' \end{bmatrix}}_{NK \times NT} \underbrace{\begin{bmatrix} Z^{(1)} & \cdots & 0_{T \times K} \\ \vdots & \ddots & \vdots \\ 0_{T \times K} & \cdots & Z^{(N)} \end{bmatrix}}_{NT \times NK}$$

²Recall that the corresponding equations obtained from the reduced form —the first row of the full set of equations in matrix form— comprises K linear equations on $N + K$ unknown structural parameters, so we need at least N additional equations.

$$= \underbrace{\begin{bmatrix} X'Z^{(1)} & \cdots & 0_{K \times K} \\ \vdots & \ddots & \vdots \\ 0_{K \times K} & \cdots & X'Z^{(N)} \end{bmatrix}}_{NK \times NK}$$

is a square matrix with square, nonsingular diagonal blocks $X'Z^{(i)}$ and is therefore nonsingular. Thus, in the case of an exactly identified system, the system-2SLS estimator simplifies to

$$\begin{aligned} \hat{\delta}_{2SLS} &= (\hat{Z}^{*'} Z^*)^{-1} \hat{Z}^{*'} Y^* \\ &= \left(\left((I_N \otimes X) (I_N \otimes (X'X)^{-1}) (I_N \otimes X') Z^* \right)' Z^* \right)^{-1} \left((I_N \otimes X) (I_N \otimes (X'X)^{-1}) (I_N \otimes X') Z^* \right)' Y^* \\ &= \left(Z^{*'} (I_N \otimes X) (I_N \otimes (X'X)^{-1}) (I_N \otimes X') Z^* \right)^{-1} Z^{*'} (I_N \otimes X) (I_N \otimes (X'X)^{-1}) (I_N \otimes X') Y^* \\ &= \left((I_N \otimes X') Z^* \right)^{-1} \left(I_N \otimes (X'X)^{-1} \right)^{-1} \left(Z^{*'} (I_N \otimes X) \right)^{-1} Z^{*'} (I_N \otimes X) (I_N \otimes (X'X)^{-1}) (I_N \otimes X') Y^* \\ &= \left((I_N \otimes X') Z^* \right)^{-1} \left(I_N \otimes (X'X)^{-1} \right)^{-1} (I_N \otimes (X'X)^{-1}) (I_N \otimes X') Y^* \\ &= \left((I_N \otimes X') Z^* \right)^{-1} (I_N \otimes X') Y^*. \end{aligned}$$

Similarly, the 3SLS estimator simplifies to

$$\begin{aligned} \hat{\delta}_{3SLS} &= \left((I_N \otimes X') Z^* \right)^{-1} \left(\Sigma^{-1} \otimes (X'X)^{-1} \right)^{-1} \left(Z^{*'} (I_N \otimes X) \right)^{-1} Z^{*'} (I_N \otimes X) (\Sigma^{-1} \otimes (X'X)^{-1}) (I_N \otimes X') Y^* \\ &= \left((I_N \otimes X') Z^* \right)^{-1} \left(\Sigma^{-1} \otimes (X'X)^{-1} \right)^{-1} (\Sigma^{-1} \otimes (X'X)^{-1}) (I_N \otimes X') Y^* \\ &= \left((I_N \otimes X') Z^* \right)^{-1} (I_N \otimes X') Y^* \\ &= \hat{\delta}_{2SLS}, \end{aligned}$$

which coincides with the system-2SLS estimator for the exactly identified system.

4 Question 2

4.1 General derivations

Consider the MLE of (δ, Σ) under the assumption that

$$u_t \stackrel{iid}{\sim} N(0, \Sigma) \implies v_t \stackrel{iid}{\sim} N(0, \Omega),$$

where $\Omega = \tilde{A}^{-1}\Sigma\tilde{A}^{-1'}$ and $v_t \equiv \tilde{A}^{-1}u_t$ is the projection error in the reduced form

$$y_t = \Pi x_t + v_t$$

with $\Pi \equiv -\tilde{A}^{-1}\tilde{B}$.³ We can obtain the matrix representation of the reduced form by transposing and stacking the reduced form for the T observations:

$$Y = X\Pi' + V,$$

where

$$V \equiv \begin{pmatrix} v'_1 \\ \vdots \\ v'_T \end{pmatrix}_{T \times N} = U\tilde{A}^{-1'}.$$

Vectorize V to obtain multivariate-normal vector⁴

$$V^* \equiv \text{vec}(V) = \begin{pmatrix} v_{11} \\ \vdots \\ v_{1T} \\ \vdots \\ v_{N1} \\ \vdots \\ v_{NT} \end{pmatrix}_{NT \times 1} \sim N(0, \Omega \otimes I_T)$$

since

$$\mathbb{E}[V^*V^{*'}] = \mathbb{E} \left[\begin{pmatrix} v_{11} \\ \vdots \\ v_{1T} \\ \vdots \\ v_{N1} \\ \vdots \\ v_{NT} \end{pmatrix} \begin{pmatrix} v_{11} & \cdots & v_{1T} & \cdots & v_{N1} & \cdots & v_{NT} \end{pmatrix} \right]$$

³Notice that the variance of v_t , Ω , is a function of the structural parameters in δ and Σ , $\Omega(\delta, \Sigma)$.

⁴An alternative is to work directly with the [matrix normal distribution](#).

$$\begin{aligned}
&= \underbrace{\begin{bmatrix} \underbrace{\mathbb{E}[v_{11}^2]}_{\omega_{11}} & \cdots & \underbrace{\mathbb{E}[v_{11}v_{1T}]}_0 & \cdots & \underbrace{\mathbb{E}[v_{11}v_{N1}]}_{\omega_{1N}} & \cdots & \underbrace{\mathbb{E}[v_{11}v_{NT}]}_0 \\ \vdots & \ddots & \vdots & \ddots & \vdots & \ddots & \vdots \\ \underbrace{\mathbb{E}[v_{1T}v_{11}]}_0 & \cdots & \underbrace{\mathbb{E}[v_{1T}^2]}_{\omega_{11}} & \cdots & \underbrace{\mathbb{E}[v_{1T}v_{N1}]}_0 & \cdots & \underbrace{\mathbb{E}[v_{1T}v_{NT}]}_{\omega_{1N}} \\ \vdots & \ddots & \vdots & \ddots & \vdots & \ddots & \vdots \\ \underbrace{\mathbb{E}[v_{N1}v_{11}]}_{\omega_{N1}} & \cdots & \underbrace{\mathbb{E}[v_{N1}v_{1T}]}_0 & \cdots & \underbrace{\mathbb{E}[v_{N1}^2]}_{\omega_{NN}} & \cdots & \underbrace{\mathbb{E}[v_{N1}v_{NT}]}_0 \\ \vdots & \ddots & \vdots & \ddots & \vdots & \ddots & \vdots \\ \underbrace{\mathbb{E}[v_{NT}v_{11}]}_0 & \cdots & \underbrace{\mathbb{E}[v_{NT}v_{1T}]}_{\omega_{N1}} & \cdots & \underbrace{\mathbb{E}[v_{NT}v_{N1}]}_0 & \cdots & \underbrace{\mathbb{E}[v_{NT}^2]}_{\omega_{NN}} \end{bmatrix}}_{NT \times NT} \\
&= \underbrace{\begin{bmatrix} \omega_{11} & \cdots & \omega_{1N} \\ \vdots & \ddots & \vdots \\ \omega_{N1} & \cdots & \omega_{NN} \end{bmatrix}}_{N \times N} \otimes \underbrace{\begin{bmatrix} 1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & 1 \end{bmatrix}}_{T \times T} \\
&= \Omega \otimes I_T.
\end{aligned}$$

Notice that

$$\begin{aligned}
V^{*'}(\Omega \otimes I_T)^{-1}V^* &= V^{*'}(\Omega^{-1} \otimes I_T^{-1})V^* \\
&= V^{*'}(\Omega^{-1'} \otimes I_T)\text{vec}(V) \\
&= V^{*'}\text{vec}(I_T V \Omega^{-1}) \\
&= \text{vec}(V)'\text{vec}(V \Omega^{-1}) \\
&= \text{tr}(V' V \Omega^{-1}) \\
&= \text{tr}(V \Omega^{-1} V') \\
&= \text{tr}((Y - X\Pi')\Omega^{-1}(Y - X\Pi')')
\end{aligned}$$

and

$$\begin{aligned}
|\Omega \otimes I_T| &= |\Omega|^T |I_T|^N \\
&= |\Omega|^T 1^N
\end{aligned}$$

$$= |\Omega|^T.$$

Therefore, the log-likelihood is

$$\begin{aligned}\ell(\delta, \Sigma) &= \ln \left((2\pi)^{-NT/2} |\Omega \otimes I_T|^{-1/2} \exp \left(-\frac{1}{2} V^{*'} (\Omega \otimes I_T)^{-1} V^* \right) \right) \\ &= \ln \left((2\pi)^{-NT/2} |\Omega|^{-T/2} \exp \left(-\frac{1}{2} \text{tr} \left((Y - X\Pi') \Omega^{-1} (Y - X\Pi')' \right) \right) \right) \\ &= -\frac{NT}{2} \ln(2\pi) - \frac{T}{2} \ln(|\Omega|) - \frac{1}{2} \text{tr} \left((Y - X\Pi') \Omega^{-1} (Y - X\Pi')' \right)\end{aligned}$$

and we can minimize the objective function

$$Q(\delta, \Sigma) = \frac{NT}{2} \ln(2\pi) + \frac{T}{2} \ln(|\Omega|) + \frac{1}{2} \text{tr} \left((Y - X\Pi') \Omega^{-1} (Y - X\Pi')' \right)$$

since

$$\begin{aligned}\arg \max_{(\delta, \Sigma)} \ell(\delta, \Sigma) &= \arg \min_{(\delta, \Sigma)} -\ell(\delta, \Sigma) \\ &= \arg \min_{(\delta, \Sigma)} Q(\delta, \Sigma).\end{aligned}$$

4.2 Recursive system

We define a recursive system as one where \tilde{A} is lower triangular and Σ is diagonal:

$$\tilde{A} = \begin{bmatrix} \tilde{a}_1^1 & 0 & 0 & \cdots & 0 \\ \tilde{a}_1^2 & \tilde{a}_2^2 & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \tilde{a}_1^N & \tilde{a}_2^N & \tilde{a}_3^N & \cdots & \tilde{a}_N^N \end{bmatrix} \quad \Sigma = \begin{bmatrix} \sigma_{11} & 0 & 0 & \cdots & 0 \\ 0 & \sigma_{22} & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & \cdots & \sigma_{NN} \end{bmatrix}.$$

The normalizations $\tilde{a}_{ii} = 1$ for $i \in \{1, \dots, N\}$ further imply that A is unitriangular and so is A^{-1} .

Moreover, A^{-1} lower unitriangular implies that $A^{-1'}$ is upper unitriangular. Therefore,

$$\begin{aligned}|\Omega| &= |A^{-1} \Sigma A^{-1'}| \\ &= \underbrace{|A^{-1}|}_1 |\Sigma| \underbrace{|A^{-1'}|}_1 \\ &= |\Sigma| \\ &= \prod_{i=1}^N \sigma_{ii}.\end{aligned}$$

Also, notice that

$$\begin{aligned}
Y - X\Pi' &= \begin{bmatrix} y_1' \\ \vdots \\ y_T' \end{bmatrix} - \begin{bmatrix} x_1' \\ \vdots \\ x_T' \end{bmatrix} \Pi' \\
&= \begin{bmatrix} y_1' - x_1' \Pi' \\ \vdots \\ y_T' - x_T' \Pi' \end{bmatrix} \\
&= \begin{bmatrix} (y_1 - \Pi x_1)' \\ \vdots \\ (y_T - \Pi x_T)' \end{bmatrix}
\end{aligned}$$

and

$$\Omega^{-1} = A' \Sigma^{-1} A$$

Therefore,

$$\begin{aligned}
(Y - X\Pi') \Omega^{-1} (Y - X\Pi')' &= \begin{bmatrix} (y_1 - \Pi x_1)' \Omega^{-1} \\ \vdots \\ (y_T - \Pi x_T)' \Omega^{-1} \end{bmatrix} \begin{bmatrix} (y_1 - \Pi x_1) & \cdots & (y_T - \Pi x_T) \end{bmatrix} \\
&= \begin{bmatrix} (y_1 - \Pi x_1)' \Omega^{-1} (y_1 - \Pi x_1) & \cdots & (y_1 - \Pi x_1)' \Omega^{-1} (y_T - \Pi x_T) \\ \vdots & \ddots & \vdots \\ (y_T - \Pi x_T)' \Omega^{-1} (y_1 - \Pi x_1) & \cdots & (y_T - \Pi x_T)' \Omega^{-1} (y_T - \Pi x_T) \end{bmatrix}
\end{aligned}$$

and

$$\begin{aligned}
\text{tr} \left((Y - X\Pi') \Omega^{-1} (Y - X\Pi')' \right) &= \sum_{t=1}^T (y_t - \Pi x_t)' \Omega^{-1} (y_t - \Pi x_t) \\
&= \sum_{t=1}^T (y_t - (-A^{-1}B)x_t)' A' \Sigma^{-1} A (y_t - (-A^{-1}B)x_t) \\
&= \sum_{t=1}^T (Ay_t + AA^{-1}Bx_t)' \Sigma^{-1} (Ay_t + AA^{-1}Bx_t) \\
&= \sum_{t=1}^T (Ay_t + Bx_t)' \Sigma^{-1} (Ay_t + Bx_t)
\end{aligned}$$

$$\begin{aligned}
&= \sum_{t=1}^T u_t' \Sigma^{-1} u_t \\
&= \sum_{t=1}^T \begin{pmatrix} y_{1t} - z_t^{(1)'} \delta^{(1)} & \cdots & y_{Nt} - z_t^{(N)'} \delta^{(N)} \end{pmatrix} \begin{bmatrix} \sigma_{11}^{-1} & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \sigma_{NN}^{-1} \end{bmatrix} \begin{pmatrix} y_{1t} - z_t^{(1)'} \delta^{(1)} \\ \vdots \\ y_{Nt} - z_t^{(N)'} \delta^{(N)} \end{pmatrix} \\
&= \sum_{t=1}^T \sum_{i=1}^N \frac{\left(y_{it} - z_t^{(i)'} \delta^{(i)} \right)^2}{\sigma_{ii}}.
\end{aligned}$$

These results imply that we can write the objective function as

$$\begin{aligned}
Q(\delta, \Sigma) &= \frac{NT}{2} \ln(2\pi) + \frac{T}{2} \ln \left(\prod_{i=1}^N \sigma_{ii} \right) + \frac{1}{2} \sum_{t=1}^T \sum_{i=1}^N \frac{\left(y_{it} - z_t^{(i)'} \delta^{(i)} \right)^2}{\sigma_{ii}} \\
&= \frac{NT}{2} \ln(2\pi) + \frac{T}{2} \sum_{i=1}^N \ln(\sigma_{ii}) + \frac{1}{2} \sum_{t=1}^T \sum_{i=1}^N \frac{\left(y_{it} - z_t^{(i)'} \delta^{(i)} \right)^2}{\sigma_{ii}} \\
&= \sum_{i=1}^N \left\{ \frac{T}{2} \ln(2\pi) + \frac{T}{2} \ln(\sigma_{ii}) + \frac{1}{2} \sum_{t=1}^T \frac{\left(y_{it} - z_t^{(i)'} \delta^{(i)} \right)^2}{\sigma_{ii}} \right\}.
\end{aligned}$$

Finally, notice that the MLE is

$$\begin{aligned}
(\hat{\delta}, \hat{\Sigma}) &= \arg \min_{(\delta, \Sigma)} Q(\delta, \Sigma) \\
&= \arg \min_{(\delta, \Sigma)} \sum_{i=1}^N \left\{ \frac{T}{2} \ln(2\pi) + \frac{T}{2} \ln(\sigma_{ii}) + \frac{1}{2} \sum_{t=1}^T \frac{\left(y_{it} - z_t^{(i)'} \delta^{(i)} \right)^2}{\sigma_{ii}} \right\},
\end{aligned}$$

which can be obtained by minimizing

$$\frac{T}{2} \ln(2\pi) + \frac{T}{2} \ln(\sigma_{ii}) + \frac{1}{2} \sum_{t=1}^T \frac{\left(y_{it} - z_t^{(i)'} \delta^{(i)} \right)^2}{\sigma_{ii}}$$

for each $i \in \{1, \dots, N\}$, i.e., OLS equation by equation (MLE based on the marginal distribution of u_{it}).